MASTER'S DEEP LEARNING

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- For example, *linear regression*.
- Linear model: consider a linear function

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$$y(\mathbf{x}, \mathbf{w}) = w_0 + \sum_{j=1}^p x_j w_j = \mathbf{x}^\top \mathbf{w}, \quad \mathbf{x} = (1, x_1, \dots, x_p).$$

• How can we find optimal parameters $\hat{\mathbf{w}}$ by training data of the form $(\mathbf{x}_i, y_i)_{i=1}^N$?

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- How can we find optimal parameters $\hat{\mathbf{w}}$ by training data of the form $(\mathbf{x}_i, y_i)_{i=1}^N$?
- Least squares estimation: we will minimize

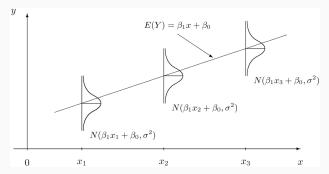
$$\mathrm{RSS}(\mathbf{w}) = \sum_{i=1}^N (y_i - \mathbf{x}_i^\top \mathbf{w})^2.$$

• There is even an exact solution, but that's not important right now.

• What is important: suppose that noise (error in the data) has a normal distribution, i.e., observed variable *t* is

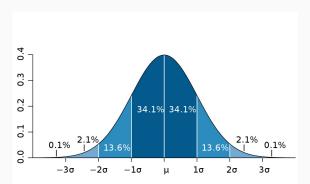
$$t = y(\mathbf{x}, \mathbf{w}) + \epsilon, \quad \epsilon \sim \mathcal{N}(0, \sigma^2), \text{ to ectb}$$

$$p(t \mid \mathbf{x}, \mathbf{w}, \sigma^2) = \mathcal{N}(t \mid y(\mathbf{x}, \mathbf{w}), \sigma^2).$$



• Aside – normal distribution:

$$p(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$



• Why is it so important?

- Consider a dataset $\mathbf{X} = \{\mathbf{x}_1, \dots, \mathbf{x}_N\}$ with correct answers $\mathbf{t} = \{t_1, \dots, t_N\}.$
- We assume that the data points are independent identically distributed:

$$p(\mathbf{t} \mid \mathbf{X}, \mathbf{w}, \sigma^2) = \prod_{n=1}^N \mathcal{N}\left(t_n \mid \mathbf{w}^\top \phi(\mathbf{x}_n), \sigma^2\right).$$

• We take the logarithm (we omit **X** below for brevity):

$$\ln p(\mathbf{t} \mid \mathbf{w}, \sigma^2) = -\frac{N}{2} \ln(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum_{n=1}^N \left(t_n - \mathbf{w}^\top \phi(\mathbf{x}_n) \right)^2.$$

• And we see that to maximize the likelihood w.r.t. **w** we need to minimze mean squared error!

$$\nabla_{\mathbf{w}} \ln p(\mathbf{t} \mid \mathbf{w}, \sigma^2) = \frac{1}{\sigma^2} \sum_{n=1}^{N} \left(t_n - \mathbf{w}^\top \phi(\mathbf{x}_n) \right) \phi(\mathbf{x}_n).$$

- We can also get a posterior distribution, introducing prior distributions (also normal).
- And then the predictive distribution

$$p(y \mid \mathbf{x}, D) = \int_{\mathbf{w}} p(y \mid \mathbf{x}, \mathbf{w}) p(\mathbf{w} \mid D) \mathrm{d}\mathbf{w}$$

...but that's beside the point right now.

• Main conclusion: in many regression problems it makes sense to minimize the sum of squared deviations, this corresponds to normally distributed noise.

BAYESIAN REGULARIZATION

- And now let us look at regression from the pure Bayesian perspective.
- Recall that in Bayesian inference, we
 - (1) find the posterior distribution на гипотезах/параметрах:

 $p(\boldsymbol{\theta} \mid \boldsymbol{D}) \propto p(\boldsymbol{D} | \boldsymbol{\theta}) p(\boldsymbol{\theta})$

(and/or find the maximal a posteriori hypothesis arg max_θp(θ | D));
(2) find the predictive distribution:

$$p(x \mid D) \propto \int_{\theta \in \Theta} p(x \mid \theta) p(D|\theta) p(\theta) \mathrm{d}\theta.$$

- \cdot We have not yet had any priors in our study of linear regression.
- Let us introduce a prior; e.g., the normal distribution:

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{w} \mid \boldsymbol{\mu}_0, \boldsymbol{\Sigma}_0).$$

• Consider a dataset $\mathbf{X} = {\mathbf{x}_1, ..., \mathbf{x}_N}$ with values $\mathbf{t} = {t_1, ..., t_N}$; we again assume that they are independent and identically distributed:

$$p(\mathbf{t} \mid \mathbf{X}, \mathbf{w}, \sigma^2) = \prod_{n=1}^N \mathcal{N}\left(t_n \mid \mathbf{w}^\top \phi(\mathbf{x}_n), \sigma^2\right).$$

 \cdot Then the problem is to compute

$$\begin{split} p(\mathbf{w} \mid \mathbf{t}) &\propto p(\mathbf{t} \mid \mathbf{X}, \mathbf{w}, \sigma^2) p(\mathbf{w}) \\ &= \mathcal{N}(\mathbf{w} \mid \boldsymbol{\mu}_0, \boldsymbol{\Sigma}_0) \prod_{n=1}^N \mathcal{N}\left(t_n \mid \mathbf{w}^\top \boldsymbol{\phi}(\mathbf{x}_n), \sigma^2\right). \end{split}$$

• Let us compute!

• We get

$$\begin{split} \boldsymbol{\wp}(\mathbf{w} \mid \mathbf{t}) &= \mathcal{N}(\mathbf{w} \mid \boldsymbol{\mu}_N, \boldsymbol{\Sigma}_N), \\ \boldsymbol{\mu}_N &= \boldsymbol{\Sigma}_N \left(\boldsymbol{\Sigma}_0^{-1} \boldsymbol{\mu}_0 + \frac{1}{\sigma^2} \boldsymbol{\Phi}^\top \mathbf{t}\right), \\ \boldsymbol{\Sigma}_N &= \left(\boldsymbol{\Sigma}_0^{-1} + \frac{1}{\sigma^2} \boldsymbol{\Phi}^\top \boldsymbol{\Phi}\right)^{-1}. \end{split}$$

• And now the log likelihood.

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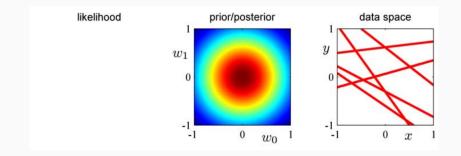
• If we take the prior distribution around zero:

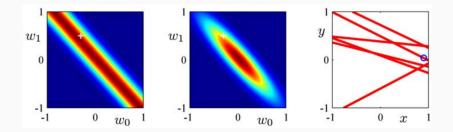
$$p(\mathbf{w}) = \mathcal{N}(\mathbf{w} \mid 0, \frac{1}{\alpha}\mathbf{I}),$$

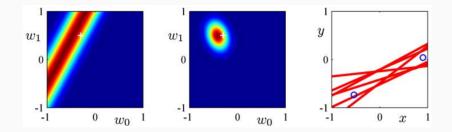
we get the log likelihood as

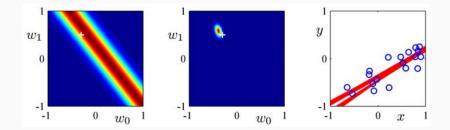
$$\ln p(\mathbf{w} \mid \mathbf{t}) = -\frac{1}{2\sigma^2} \sum_{n=1}^{N} \left(t_n - \mathbf{w}^\top \phi(\mathbf{x}_n) \right)^2 - \frac{\alpha}{2} \mathbf{w}^\top \mathbf{w} + \text{const},$$

i.e., precisely ridge regression!









• A slight generalization – a more general prior distribution:

$$p(\mathbf{w} \mid \alpha) = \left[\frac{q}{2} \left(\frac{\alpha}{2}\right)^{1/q} \frac{1}{\Gamma(1/q)}\right]^M e^{-\frac{\alpha}{2} \sum_{j=1}^M \left|w_j\right|^q}$$

Упражнение. Compute the log likelihood.

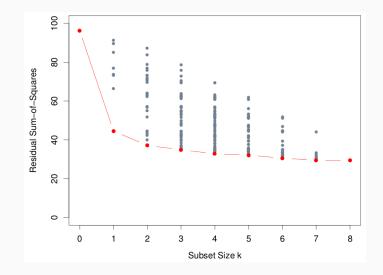
REGULARIZATION AGAIN

- We know that least squares do not always work well. Two reasons:
 - bad predictive power often better to regularize, trading bias for variance;
 - 2. hard to interpret we often want to understand what is going on, and if we have lots of different nonzero numbers, it's hard.
- Hence, we'd like to get more nonzero components in the vector
 w.

- What if we do it directly? Simply presume most coefficients are zero and find the nonzero ones.
- This is called subset selection.
- Best subset selection: choose the subset of \boldsymbol{k} input variables that gives the best results

- Naturally, this does not work computationally: there are lots of subsets.
- Forward-stepwise selection: start from the intercept, then add one best predictor per step.
- Backward-stepwise selection: start from full regression and remove the predictor that influences the error the least.

SUBSET SELECTION



• Let us now consider lasso regression:

$$L(\mathbf{w}) = \frac{1}{2} \sum_{i=1}^{N} (f(x_i, \mathbf{w}) - y_i)^2 + \lambda \sum_{j=0}^{p} |w_j|.$$

- The main difference is that the form of the constraints is now such that it is much more probable to get strictly zero w_j .
- Btw, what do I mean by "form of the constraints"?

LASSO

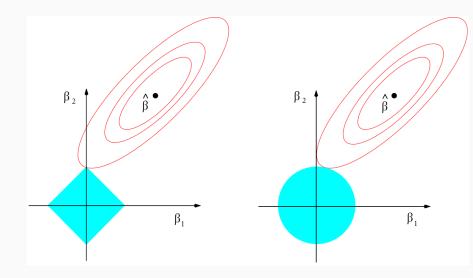
• We can rewrite the regression with regularizer in a different way:

$$\mathbf{w}^* = \arg\min_{\mathbf{w}} \left\{ \frac{1}{2} \sum_{i=1}^N (f(x_i, \mathbf{w}) - y_i)^2 + \lambda \sum_{j=0}^p |w_j| \right\},$$

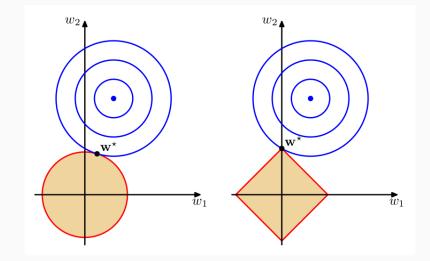
is equivalent to

$$\mathbf{w}^* = \arg\min_{\mathbf{w}} \left\{ \frac{1}{2} \sum_{i=1}^N (f(x_i, \mathbf{w}) - y_i)^2 \right\} \text{ for } \sum_{j=0}^p |w_j| \leq t.$$

Упражнение. Prove it.



RIDGE AND LASSO

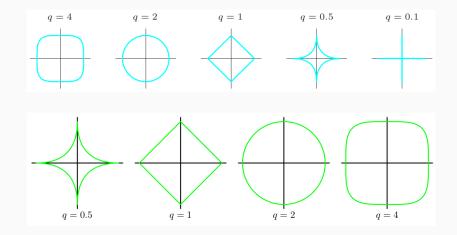


 \cdot We can generalize ridge and lasso regression to

$$L(\mathbf{w}) = \frac{1}{2} \sum_{i=1}^{N} (f(x_i, \mathbf{w}) - y_i)^2 + \lambda \sum_{j=0}^{p} (|w_j|)^q.$$

Упражнение. Which prior distribution on ${f w}$ does this correspond to?

DIFFERENT q



Thank you for your attention!