

CLUSTERING AND NAIVE BAYES

NATURAL LANGUAGE PROCESSING

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GENERALIZING NAIVE BAYES

- Naive Bayes makes two additional simplifying assumptions:
 - we know class labels for all documents;
 - each document belongs to only one class (has only one topic).
- It turns out we can remove both of these constraints.
- First, what do we do if we don't know class labels, i.e., the dataset is unlabeled?
- What kind of a problem is it and how can we solve it?
- It's clustering, but we need to start with the EM algorithm...

- Often the data has *latent* (missing) variables.
- We have the result of sampling a distribution, but some of the parameters are not known.
- We can treat latent variables as random values and look for the maximal likelihood hypothesis h , i.e., maximize

$$\mathbb{E}[p(D|h)] = \mathbb{E}\left[\int p(D, z|h)dz\right]$$

for latent variables z .

- Example: consider a random variable x sampled from a mixture of two Gaussians with the same variance σ^2 and different means μ_1, μ_2 .
- Two-stage sampling, but we don't know the first stage results.
- One point is a triple $\langle x_i, z_{i1}, z_{i2} \rangle$, where $z_{ij} = 1$ iff x_i was generated from distribution j , and we don't know z_{ij} .

- EM algorithm idea:
 - generate a hypothesis $h = (\mu_1, \mu_2)$;
 - while we have not reached local maximum:
 - compute the expectation $E(z_{ij})$ given the current hypothesis (E -step);
 - compute the new hypothesis $h' = (\mu'_1, \mu'_2)$ assuming that z_{ij} take values $E(z_{ij})$ computed before (M -step).

- For the Gaussians:

$$\begin{aligned} E(z_{ij}) &= \frac{p(x = x_i | \mu = \mu_j)}{p(x = x_i | \mu = \mu_1) + p(x = x_i | \mu = \mu_2)} = \\ &= \frac{e^{-\frac{1}{2\sigma^2}(x_i - \mu_j)^2}}{e^{-\frac{1}{2\sigma^2}(x_i - \mu_1)^2} + e^{-\frac{1}{2\sigma^2}(x_i - \mu_2)^2}}. \end{aligned}$$

- We compute the expectations and then tune the hypothesis:

$$\mu_j \leftarrow \frac{1}{m} \sum_{i=1}^m E(z_{ij}) x_i.$$

- Formally, we are maximizing the likelihood with data $\mathcal{X} = \{x_1, \dots, x_N\}$.

$$L(\theta | \mathcal{X}) = p(\mathcal{X} | \theta) = \prod p(x_i | \theta)$$

or, which is the same, maximizing $\ell(\theta | \mathcal{X}) = \log L(\theta | \mathcal{X})$.

- EM can help if this maximum is hard to find, but easy once we know something else...

- Suppose that the data has *latent variables* such that the problem would be easy if we knew them.
- They don't necessarily have to correspond to anything interesting, maybe they are there just for convenience.
- In any case, we get a dataset $\mathcal{Z} = (\mathcal{X}, \mathcal{Y})$ with joint density

$$p(z | \theta) = p(x, y | \theta) = p(y | x, \theta)p(x | \theta).$$

- Full likelihood $L(\theta | \mathcal{Z}) = p(\mathcal{X}, \mathcal{Y} | \theta)$ is a random variable since we don't know \mathcal{Y} .

- Note that the real likelihood is $L(\theta) = E_Y [p(\mathcal{X}, \mathcal{Y} | \theta) | \mathcal{X}, \theta]$.
- E-step computes the conditional expectation of the (log) full likelihood given \mathcal{X} and current estimates for parameters θ_n :

$$Q(\theta, \theta_n) = E [\log p(\mathcal{X}, \mathcal{Y} | \theta) | \mathcal{X}, \theta_n].$$

- Here θ_n are current estimates, θ are unknown values (which we want to get at the end); i.e., $Q(\theta, \theta_n)$ is a function of θ .

- E-step computes the conditional expectation of the (log) full likelihood given \mathcal{X} and current estimates for parameters θ :

$$Q(\theta, \theta_n) = E[\log p(\mathcal{X}, \mathcal{Y} | \theta) | \mathcal{X}, \theta_n].$$

- Conditional expectation:

$$E[\log p(\mathcal{X}, \mathcal{Y} | \theta) | \mathcal{X}, \theta_n] = \int_y \log p(\mathcal{X}, y | \theta) p(y | \mathcal{X}, \theta_n) dy,$$

where $p(y | \mathcal{X}, \theta_n)$ is the marginal distribution of latent variables.

- EM works best when it's easy to compute, maybe even analytically.
- Instead of $p(y | \mathcal{X}, \theta_n)$ we can substitute $p(y, \mathcal{X} | \theta_n) = p(y | \mathcal{X}, \theta_n)p(\mathcal{X} | \theta_n)$, it won't change anything.

- As a result, after the E-step of the EM algorithm we get the function $Q(\theta, \theta_n)$.
- On the M-step, we maximize

$$\theta_{n+1} = \arg \max_{\theta} Q(\theta, \theta_n).$$

- And repeat until convergence.
- Actually, it suffices to find θ_{n+1} such that $Q(\theta_{n+1}, \theta_n) > Q(\theta_n, \theta_n)$
 - Generalized EM.
- It remains to see what $Q(\theta, \theta_n)$ means and why it all works.

- We wanted to pass from θ_n to θ such that $\ell(\theta) > \ell(\theta_n)$.

$$\begin{aligned}\ell(\theta) - \ell(\theta_n) &= \\ &= \log \left(\int_y p(\mathcal{X} | y, \theta) p(y | \theta) dy \right) - \log p(\mathcal{X} | \theta_n) = \\ &= \log \left(\int_y p(y | \mathcal{X}, \theta_n) \frac{p(\mathcal{X} | y, \theta) p(y | \theta)}{p(y | \mathcal{X}, \theta_n)} dy \right) - \log p(\mathcal{X} | \theta_n) \geq \\ &\geq \int_y p(y | \mathcal{X}, \theta_n) \log \left(\frac{p(\mathcal{X} | y, \theta) p(y | \theta)}{p(y | \mathcal{X}, \theta_n)} \right) dy - \log p(\mathcal{X} | \theta_n) = \\ &= \int_y p(y | \mathcal{X}, \theta_n) \log \left(\frac{p(\mathcal{X} | y, \theta) p(y | \theta)}{p(\mathcal{X} | \theta_n) p(y | \mathcal{X}, \theta_n)} \right) dy.\end{aligned}$$

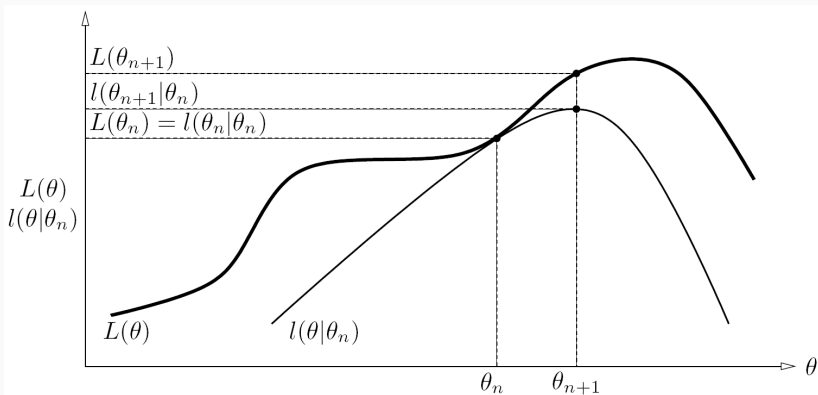
- Thus, we get

$$\begin{aligned}\ell(\theta) &\geq \ell(\theta, \theta_n) = \\ &= \ell(\theta_n) + \int_y p(y | \mathcal{X}, \theta_n) \log \left(\frac{p(\mathcal{X} | y, \theta)p(y | \theta)}{p(\mathcal{X} | \theta_n)p(y | \mathcal{X}, \theta_n)} \right) dy.\end{aligned}$$

Упражнение. Prove that $\ell(\theta_n, \theta_n) = \ell(\theta_n)$.

- In other words, we have found a lower bound on $\ell(\theta)$ everywhere that touches it at point θ_n .
- I.e., we have found a lower bound for the likelihood and move to a point that maximizes it (or at least improves).
- This is called minorization-maximization (*MM*).

JUSTIFICATION OF EM



- It remains to see that we can maximize Q .

$$\begin{aligned}
 \theta_{n+1} &= \arg \max_{\theta} \ell(\theta, \theta_n) = \arg \max_{\theta} \left\{ \ell(\theta_n) + \right. \\
 &\quad \left. + \int_y f(y | \mathcal{X}, \theta_n) \log \left(\frac{p(\mathcal{X} | y, \theta) f(y | \theta)}{p(\mathcal{X} | \theta_n) f(y | \mathcal{X}, \theta_n)} \right) dy \right\} = \\
 &= \arg \max_{\theta} \left\{ \int_y p(y | \mathcal{X}, \theta_n) \log (p(\mathcal{X} | y, \theta) p(y | \theta)) dy \right\} = \\
 &= \arg \max_{\theta} \left\{ \int_y p(y | \mathcal{X}, \theta_n) \log p(\mathcal{X}, y | \theta) dy \right\} = \\
 &= \arg \max_{\theta} \{Q(\theta, \theta_n)\},
 \end{aligned}$$

and the rest does not depend on θ .

EM AND CLUSTERING

- How can we apply EM to clustering?

- Hypothesis: test examples are drawn independently from a mixture of cluster distributions

$$p(x) = \sum_{c \in \mathcal{C}} w_c p_c(x), \quad \sum_{c \in \mathcal{C}} w_c = 1,$$

where w_c is the probability to get a point from cluster c , p_c is the density of cluster c .

- What would be the form of p_c ?

HYPOTHESIS CONT'D

- What would be the form of p_c ?
- Let's try... mmm... well, Gaussians. :)
- *Hypothesis 2*: each cluster c is a d -dimensional Gaussian distribution with mean $\mu_c = \{\mu_{c1}, \dots, \mu_{cd}\}$ and diagonal matrix of covariances $\Sigma_c = \text{diag}(\sigma_{c1}^2, \dots, \sigma_{cd}^2)$ (i.e., separate variance for every independent coordinate).

- Thus, we have formalized clustering as learning a mixture of distributions. That's where EM comes into play.
- Each test example looks like $(f_1(x), \dots, f_n(x))$.
- Latent variables in this case are probabilities g_{ic} of x_i to belong to cluster $c \in C$.

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$$g_{ic} = \frac{w_c p_c(x_i)}{\sum_{c' \in C} w_{c'} p_{c'}(x_i)}.$$

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- *M*-step: with g_{ic} we refine cluster parameters w, μ, σ :

$$w_c = \frac{1}{n} \sum_{i=1}^n g_{ic},$$

- *E*-step: by Bayes theorem, we compute latent variables g_{ic} :

$$g_{ic} = \frac{w_c p_c(x_i)}{\sum_{c' \in C} w_{c'} p_{c'}(x_i)}.$$

- *M*-step: with g_{ic} we refine cluster parameters w, μ, σ :

$$w_c = \frac{1}{n} \sum_{i=1}^n g_{ic}, \quad \mu_{cj} = \frac{1}{nw_c} \sum_{i=1}^n g_{ic} f_j(x_i),$$

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$$w_c = \frac{1}{n} \sum_{i=1}^n g_{ic}, \quad \mu_{cj} = \frac{1}{nw_c} \sum_{i=1}^n g_{ic} f_j(x_i),$$

$$\sigma_{cj}^2 = \frac{1}{nw_c} \sum_{i=1}^n g_{ic} (f_j(x_i) - \mu_{cj})^2.$$

EMCluster($X, |C|$):

- Initialize $|C|$ clusters; initial approximation: $w_c := 1/|C|$, $\mu_c := \text{random } x_i$, $\sigma_{cj}^2 := \frac{1}{n|C|} \sum_{i=1}^n (f_j(x_i) - \mu_{cj})^2$.
- While cluster composition changes:
 - E -step: $g_{ic} := \frac{w_c p_c(x_i)}{\sum_{c' \in C} w_{c'} p_{c'}(x_i)}$.
 - M -step: $w_c = \frac{1}{n} \sum_{i=1}^n g_{ic}$, $\mu_{cj} = \frac{1}{nw_c} \sum_{i=1}^n g_{ic} f_j(x_i)$,

$$\sigma_{cj}^2 = \frac{1}{nw_c} \sum_{i=1}^n g_{ic} (f_j(x_i) - \mu_{cj})^2.$$

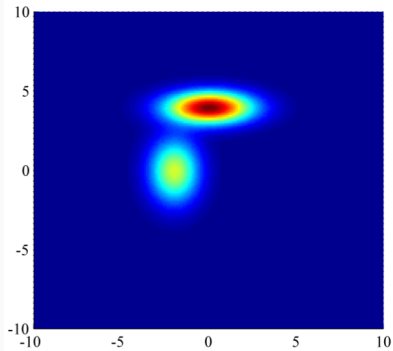
- Find which cluster x_i falls into:

$$\text{clust}_i := \arg \max_{c \in C} g_{ic}.$$

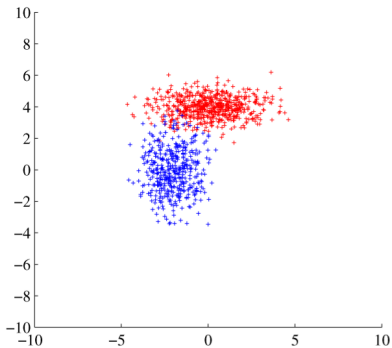
Упражнение. Prove that E-step and M-step indeed look like this.

EXAMPLE

True GMM density



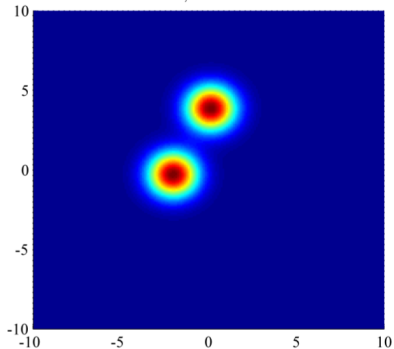
1000 i.i.d. samples



EXAMPLE

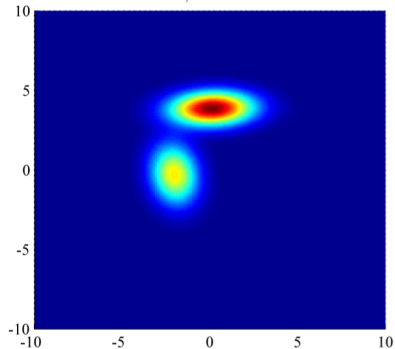
Initial Guess

$$m = 0, L^{(0)} = -3.9756$$



1st EM estimate

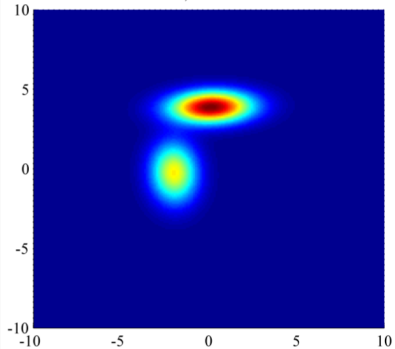
$$m = 1, L^{(1)} = -3.6492$$



EXAMPLE

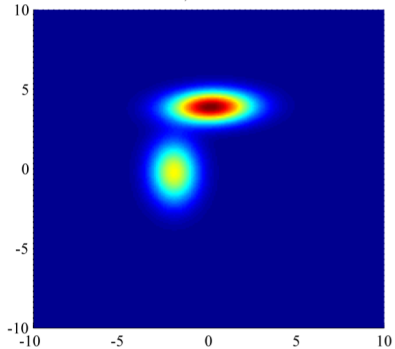
2nd EM estimate

$$m = 2, L^{(2)} = -3.6446$$



3rd EM estimate

$$m = 3, L^{(3)} = -3.6438$$



- We still need to specify the number of clusters.
- Possible solution: BIC.
- Other possible solution: non-parametric methods (out of our scope for now).

- k -means is a simplification of EM.
- Instead of computing probabilities of clusters, we use hard clustering.
- Besides, we cannot change the form of clusters in k -means (and that's not so bad).

- Formally, k -means minimizes the error

$$E(X, C) = \sum_{i=1}^n \|x_i - \mu_i\|^2,$$

where μ_i is the cluster centroid nearest to x_i .

- I.e., we move centers and automatically relate points to nearest clusters.

- So now we can come back to clustering with the naive Bayes assumptions.
- It is discrete, and it's best to solve directly with the EM algorithm (Expectation–Maximization):
 - on the E-step we compute expectations of which document belongs to which topic/class;
 - on the M-step, recompute probabilities $p(w | t)$ for fixed labels with naive Bayes.
- This is a simple generalization.
- Harder problem – generalize to multiple topics per document.

Thank you for your attention!